#### Analiza danych

## Mariusz Przybycień

Wydział Fizyki i Informatyki Stosowanej Akademia Górniczo-Hutnicza

Wykład 13

## The notion of a correlation function

Lecture based on: W. Kittel, E.A. De Wolf, Soft Multihadron Dynamics, WS 2005 and C.A. Pruneau, Data Analysis Techniques for Physical Sciences, CUP 2017

• Consider a measurement of the numbers of particles  $N_i$  produced in volumes  $\Omega_i$ "centered" around points  $\vec{p}_i$ , (i = 1, 2) in momentum space:



 $p_{\mathrm{T},i}^{\min} \leq p_{\mathrm{T},i} \leq p_{\mathrm{T},i}^{\max}$  $\eta_i^{\min} \leq \eta_i \leq \eta_i^{\max}$  $\phi_i^{\min} \leq \phi_i \leq \phi_i^{\max}$ 

• Given the stochastic nature of particle production, the yields  $N_i$  are expected to fluctuate around the means:

$$\langle N_i \rangle = \int_{\Omega_i} \frac{d^3 N_i}{dp_{\rm T} d\phi d\eta} \, dp_{\rm T} d\phi d\eta$$

- Fluctuations about the mean are usually characterized by variance or covariance.
- Correlation function is defined as the scaled covariance in the limit in which bin sizes  $\Omega_1$  and  $\Omega_2$  vanish:

$$C(\vec{p}_1, \vec{p}_2) = \frac{1}{\Omega_1 \Omega_2} \left[ \langle N(\vec{p}_1) N(\vec{p}_2) \rangle - \langle N(\vec{p}_1) \rangle \langle N(\vec{p}_2) \rangle \right]$$



## Single- and two-particle densities

• For finite bin sizes, the ratios provide estimators of the single,  $\rho_1(\vec{p_i})$ , and the joint two-particle,  $\rho_2(\vec{p_1}, \vec{p_2})$ , density functions:

$$\hat{\rho}_1(\vec{p}_i) = \frac{\langle N(\vec{p}_i) \rangle}{\Omega_i} \xrightarrow[\Omega_i \to 0]{} \rho_1(\vec{p}_i) = \frac{d^3 N(\vec{p}_i)}{dp_{\rm T} d\phi d\eta}$$
$$\hat{\rho}_2(\vec{p}_1, \vec{p}_2) = \frac{\langle N(\vec{p}_1) N(\vec{p}_2) \rangle}{\Omega_1 \Omega_2} \xrightarrow[\Omega_{1,2} \to 0]{} \rho_2(\vec{p}_1, \vec{p}_2) = \frac{d^6 N_{\rm pairs}(\vec{p}_1, \vec{p}_2)}{dp_{\rm T,1} d\phi_1 d\eta_1 dp_{\rm T,2} d\phi_2 d\eta_2}$$

• Thus, the two-particle correlation function can be expressed in terms of density functions as:

 $C(\vec{p}_1, \vec{p}_2) = \rho_2(\vec{p}_1, \vec{p}_2) - \rho_1(\vec{p}_1)\rho_1(\vec{p}_2)$ 

- In its most general form, the two-particle correlation function  $C(\vec{p_1},\vec{p_2})$  is a function of six coordinates.
- It can be positive, null or negative (as the covariance).
- However, a measurement of correlation function can be reduced to a smalller number of coordinates of interest by integrating (marginalization) over variable that are not of interest.



#### Multiparticle Densities and Factorial Moments

- Let  $y \equiv \{p_x, p_y, p_z, p_T, \eta, \phi, ...\}$  denote all particle kinematic variables under interest in a particular study. Then, the joint-probability distribution function for n particles of the same species can be written as  $P_n(y_1, y_2, y_3, ..., y_n)$
- The differential desities  $\rho_n(y_1,...,y_n)$  are proportional to the joint probabilities:

$$\rho_n(y_1, ..., y_n) = \langle N(N-1)...(N-n+1) \rangle P_n(y_1, y_2, y_3, ..., y_n)$$

• Integration of densities over the moemntum volume  $\Omega$ , thus yields the following important relations:

$$\begin{split} \int_{\Omega} \rho_1(y) dy &= \int_{\Omega} \frac{d^3 N_i}{dp_{\rm T} d\phi d\eta} dp_{\rm T} d\phi d\eta = \langle N \rangle \\ \iint_{\Omega} \rho_2(y_1, y_2) dy_1 dy_2 &= \langle N(N-1) \rangle \\ & \dots \\ \int \dots \int_{\Omega} \rho_2(y_1, \dots, y_n) dy_1 \dots dy_n &= \langle N(N-1) \dots (N-n+1) \rangle \end{split}$$

• The averages  $\langle N(N-1)...(N-n+1) \rangle$  are called factorial moments of order n.

## Cumulants

- Inclusive *n*-particle densities  $\rho_n(y_1, ..., y_n)$  are the result of a superposition, in general, of several subprocesses (even from *n* distinct and uncorrelated subprocesses!).
- Measured *n*-tuples of particles may then feature a broad variety of correlation sources associated with a plurality of dynamic processes.
- It is a common goal of multiparticle production measurements to identify and study these correlated emissions as distinct subprocesses.
- This can be accomplished by invoking correlation functions known as (factorial) cumulans, expressed either in terms of integral correlators or as differential functions of one or more particle coordinates.
- Digression (statistical independence in terms of particle densities):

Two variables are said to be statistically independent iff their joint probability density factorizes.

The statistical independence for two particles means  $\rho_2(y_1, y_2) = \rho_1(y_1)\rho_1(y_2)$ Similarly for *n* particles we have  $\rho_n(y_1, ..., y_n) = \rho_1(y_1)...\rho_1(y_n)$ 

## Cumulants

- Cumulants of order m,  $C_m$ , are defined as m-particle densities representing emission (production) of m correlated particles originating from a common process.
- An *n*-particle density can then be expressed as a sum of several terms yielding *n* particles, but each with its own "cluster" decomposition into products of *m*-cumulants:



## Particle denstities and cumulants

• In general *n*-particle densities can be expressed in terms of cumulants using the formula (shorthand notation  $y_i \rightarrow i$ ):

$$C_n(1,...,n) = C_n(1,...,n) + \sum_{\text{perm}} C_1(1)C_{n-1}(2,...,n) + \sum_{\text{perm}} C_1(1)C_1(2)C_{n-2}(3,...,n) + \sum_{\text{perm}} C_2(1,2)C_{n-2}(3,...,n) + ... + \prod_{i=1}^n C_1(i)$$

- *m*-cumulants represent fractions of the particle production cross-section associated with processes yielding *m* correlated particles (which cannot be further factorized).
- *m*-cumulants are directly calculated based on theoretical models:



• Experimentally measured quantities are *n*-particle densities, not cumulants.

ρ

## Cumulants in terms of particle densities

• Cumulants can be obtained from measured densities using "reverse engineering":



## Cumulants scaling with source multiplicity

- Cumulants  $C_n(y_1, ..., y_n)$  feature a simple scaling property for collision systems consisting of a superposition of  $m_s$  independent (but otherwise identical) sources.
- Example: Heavy ion collisions (A+A) can be regarded (to first approximation) as a superposition of  $m_s$  nucleon-nucleon (pp) interactions, each of which produces clusters consisting of n correlated particles.

Assume that production of such clusters in pp may be described by cumulant  $C_n^{pp}$ .

At a given impact parameter b (centrality), A+A collisions should involve an average of  $\langle m_s\rangle~pp$  interactions.

 $m_s$  fluctuates from event to event, but the  $n\mbox{-}{\rm cumulant}$  for A+A collisions, at fixed  $m_s$  may be written as:

#### $C_n^{\rm AA}(y_1,...,y_n\,|\,m_s)=m_sC_n^{pp}(y_1,...,y_n)$

Averaging over all A+A collisions (and assuming a superposition of independent and unmodified pp collisions, and such that produced particles do not interact with one another) yields:

$$C_n^{\rm AA}(y_1,...,y_n) = \langle m_s \rangle C_n^{pp}(y_1,...,y_n)$$

• The total multiplicity of produced particles in A+A also features the same scaling with  $m_s$ :  $\rho_1^{AA}(y) = m_s \rho_1^{pp}(y) \Rightarrow \langle n \rangle_{AA} = m_s \langle n \rangle_{pp}$ 

#### Cumulants scaling with source multiplicity

• For the pairs of particles, one can form pairs from each of  $m_s$  individual pp collisions, but one can also mix particles from independent sources:

 $\rho_2^{\mathsf{AA}}(y_1, y_2) = m_s \rho_2^{pp}(y_1, y_2) + m_s(m_s - 1)\rho_1^{pp}(y_1)\rho_1^{pp}(y_2)$ 

• The same result can be obtained using the cumulant decomposition of

$$\begin{split} \rho_2^{AA}(y_1, y_2) &= C_1^{AA}(y_1) C_1^{AA}(y_2) + C_2^{AA}(y_1, y_2) \\ &= m_s^2 C_1^{pp}(y_1) C_1^{pp}(y_2) + m_s C_2^{pp}(y_1, y_2) \\ &= m_s^2 \rho_1^{pp}(y_1) \rho_1^{pp}(y_2) + m_s \left[ \rho_2^{pp}(y_1, y_2) - \rho_1^{pp}(y_1) \rho_1^{pp}(y_2) \right] \\ &= m_s (m_s - 1) \rho_1^{pp}(y_1) \rho_1^{pp}(y_2) + m_s \rho_2^{pp}(y_1, y_2) \end{split}$$

• At fixed value of  $m_s$ , integration over  $y_1$  and  $y_2$  yields:

$$\langle n(n-1)\rangle_{\mathsf{AA}} = m_s \langle n(n-1)\rangle_{pp} + m_s (m_s-1) \langle n \rangle_{pp}^2$$

For large  $m_s$ , the scaling is dominated by uncorrelted combinatorial pairs from particles produced in different pp interactions and approximatelly scales by  $m_s^2$ .

## Cumulants scaling with source multiplicity

• Similarly, in the case of triplets, one can show that:

$$\begin{split} \rho_{3}^{\text{AA}}(1,2,3) &= C_{1}^{\text{AA}}(1)C_{1}^{\text{AA}}(2)C_{1}^{\text{AA}}(3) + C_{1}^{\text{AA}}(1)C_{2}^{\text{AA}}(2,3) + \\ &+ C_{1}^{\text{AA}}(2)C_{2}^{\text{AA}}(1,3) + C_{1}^{\text{AA}}(3)C_{2}^{\text{AA}}(1,2) + C_{3}^{\text{AA}}(1,2,3) \\ &= m_{s}^{3}C_{1}^{pp}(1)C_{1}^{pp}(2)C_{1}^{pp}(3) + m_{s}^{2}\sum_{\text{perm}}C_{1}^{pp}(1)C_{2}^{pp}(2,3) + m_{s}C_{3}^{pp}(1,2,3) \\ &= (m_{s}^{3} - m_{s}^{2} + 2m_{s})\rho_{1}^{pp}(1)\rho_{1}^{pp}(2)\rho_{1}^{pp}(3) \\ &+ (m_{s}^{2} - m_{s})\sum_{\text{perm}}\rho_{1}^{pp}(1)\rho_{2}^{pp}(2,3) + m_{s}\rho_{3}^{pp}(1,2,3) \end{split}$$

• At fixed  $m_s$ , after integration over coordinates  $y_1, y_2$  and  $y_3$  one gets:

$$\langle n(n-1)(n-2) \rangle_{\mathsf{A}\mathsf{A}} = (m_s^3 - m_s^2 + 2m_s) \langle n \rangle_{pp}^3 + 3(m_s^2 - m_s) \langle n(n-1) \rangle_{pp} \langle n \rangle_{pp} + m_s \langle n(n-1)(n-2) \rangle_{pp}$$

The average number of triplets in A+A collisions is dominated by combinatorics and essentially scales as  $m_s^3 \langle n \rangle_{pp}^3$ .

• By extension, we conclude that the average number of *n*-ntuplets in A+A collisions scales as  $m_s^n \langle n \rangle_{pp}^n$ .

# Normalized cumulants and normalized factorial moments

• Normalized inclusive densities and normalized cumulants are defined as:

$$r_n(y_1,...,y_n) = \frac{\rho_n(y_1,...,y_n)}{\rho_1(y_1)...\rho_1(y_n)} \qquad R_n(y_1,...,y_n) = \frac{C_n(y_1,...,y_n)}{\rho_1(y_1)...\rho_1(y_n)}$$

• It is also common to use reduced (normalized) factorial moments:

$$f_n = \frac{\langle N(N-1)...(N-n+1)\rangle}{\langle N \rangle^n}$$

• For systems consisting of m identical subprocesses the normalized n-cumulant scales inversely as  $m^{n-1}$  times the n-cumulant of the subsystem ( $R_n^{(m)}$  are diluted by power  $m^{n-1}$  relative to the subsystems'  $R_n^{(1)}$ ):

$$R_n^{(m)}(y_1, ..., y_n) = \frac{C_n^{(m)}(y_1, ..., y_n)}{\rho_1^{(m)}(y_1) ... \rho_1^{(m)}(y_n)} = \frac{1}{m^{n-1}} R_n^{(1)}(y_1, ..., y_n)$$

• A simple relationship exists between the normalized desities and cumulants:  $r_2(1,2) = 1 + R_2(1,2)$  $r_3(1,2,3) = 1 + \sum_{(3)} R_2(1,2) + R_3(1,2,3)$ 

$$\begin{array}{c} & (1,2,3,4) = 1 + \sum_{(6)} R_2(1,2) + \sum_{(3)} R_2(1,2) R_2(3,4) + \sum_{(3)} R_3(1,2,3) + R_4(1,2,3,4) \\ & (3) \end{array}$$

#### Particle probability densities

• Particle probability densities have been defined as (slide 3):

$$P_n(y_1, ..., y_n) \equiv \frac{\rho_n(y_1, ..., y_n)}{\langle N(N-1)...(N-n+1) \rangle}$$

 $\bullet\,$  If the production of particles 1 to n is statistically independent, then the ratio:

$$q_n(y_1, ..., y_n) \equiv \frac{P_n(y_1, ..., y_n)}{P_1(y_1) ... P_1(y_n)} = 1$$

• From the above one can see that:

$$r_n(y_1,...,y_n) = \underbrace{\frac{\langle N(N-1)...(N-n+1)\rangle}{\langle N\rangle^n}}_{\text{multiplicity fluctuations if } \neq 1} \underbrace{q_n(y_1,...,y_n)}_{\text{genuine correlations if } \neq 1}$$

•  $q_n \neq 1$  is required to yield nonvanishig normalized cumulants, e.g.:

$$R_2(y_1, y_2) = \frac{\langle N(N-1) \rangle}{\langle N \rangle^2} q_2(y_1, y_2) - 1$$

The strength of two-particle correlations is thus determined both by the function  $q_2(y_1, y_2)$  and the amplitude of multiplicity fluctuations,  $\langle N(N-1)\rangle/\langle N\rangle^2 \neq 1$ .

## Factorial and cumulant moment-generating functions

• It is known, that for the moment generating function we have:

$$M_{\mathbf{X}}(t) = \mathcal{E}\left[e^{t\mathbf{X}}\right] = \int_{-\infty}^{\infty} e^{tx} f(x) dx \quad \Rightarrow \quad m_k = \left. \frac{d^k}{dt^k} \mathcal{E}\left[e^{t\mathbf{X}}\right] \right|_{t=0} = \left. \frac{d^k}{dt^k} M_{\mathbf{X}}(t) \right|_{t=0}$$

Inclusive densities of order n may be written as:

$$\rho_n(y_1,...,y_n) = \sum_m P_m \rho_n^{(m)}(y_1,...,y_n) \quad \Leftarrow \quad P_m \equiv \frac{\sigma_m}{\sum_m \sigma_m} = \frac{\sigma_m}{\sigma_{\text{inel}}}$$

where  $\sigma_m$  is the cross section for a the process yielding m particles, and  $\rho_n^{(m)}(...)$  are n-particle densities for processes that produce exactly m particles ( $m \ge n$ ).

• Integration of inclusive *n*-particle density yields:

$$\tilde{F}_n \equiv \int_{\Omega} \rho_n(y_1, ..., y_n) dy_1 ... dy_n = \sum_m P_m \int_{\Omega} \rho_n^{(m)}(y_1, ..., y_n) dy_1 ... dy_n$$
$$= \sum_m P_m m(m-1) ... (m-n+1) = \langle m(m-1) ... (m-n+1) \rangle \equiv \langle m^{[n]} \rangle$$

• Assuming there is a value n = N beyond which all probabilities vanish and since terms in  $P_{n < N}$  cannot contribute to  $\tilde{F}_N$  one can write:  $P_N = \tilde{F}_N / N!$ 

• Proceeding recusively, one finds:  $P_n = \frac{1}{n!} \sum_{k=0}^{N-n} (-1)^k \frac{\tilde{F}_{k+n}}{k!}$ , for n = 0, 1, ..., N

## Factorial cumulants

• The factorial moment generating function should have the form:

$$G(z) = 1 + \sum_{n=1}^{\infty} \frac{z^n}{n!} \tilde{F}_n = 1 + \sum_{n=1}^{\infty} \frac{z^n}{n!} \int_{\Omega} \rho_n(y_1, ..., y_n) dy_1 ... dy_n, \quad \tilde{F}_n = \left. \frac{d^n G(z)}{dz^n} \right|_{z=0}$$
• Factorial cumulants are defined as: 
$$f_n = \int_{\Omega} dy_1 ... \int_{\Omega} dy_n C_n(y_1, ..., y_n)$$

• Factorial moments may be expressed in terms of factorial cumulants:  $\tilde{F}_1 = f_1$   $\tilde{F}_2 = f_2 + f_1^2$   $\tilde{F}_3 = f_3 + 3f_2f_1 + f_1^3$   $\tilde{F}_4 = f_4 + 4f_3f_1 + 3f_2^2 + 6f_2f_1^2 + f_1^4$   $\tilde{F}_5 = f_5 + 5f_4f_1 + 10f_3f_2 + 10f_3f_1^2 + 15f_2^2f_1 + 10f_2f_1^3 + f_1^5$ ...

$$\tilde{F}_n = n! \sum_{\{l_i\}_n} \prod_{j=1}^n \left(\frac{f_j}{j!}\right)^{i_j} \frac{1}{l_j!}$$

where summation is done over permutations satisfying  $\sum_{i=1}^{n} il_i = n$ .

## Factorial cumulants

• Factorial cumulant generating functions are defined as:

$$\ln G(z) = \langle n \rangle z + \sum_{k=2}^{\infty} \frac{z^k}{k!} f_k \quad \Rightarrow \quad f_n = \left. \frac{d^n \ln G(z)}{dz^n} \right|_{z=0}$$

• Example: Generating function for a Poisson distribution  $P_n = \frac{\langle n \rangle^n}{n!} e^{-\langle n \rangle}$ :

$$\Rightarrow \quad G(z) = \sum_{n=0}^{\infty} P_n (1+z)^n = e^{-\langle n \rangle} \sum_{n=0}^{\infty} \frac{\langle n \rangle^n}{n!} (1+z)^n = \exp\left(\langle n \rangle z\right)$$

$$\Rightarrow \quad \tilde{F}_m = \left. \frac{d^n G(z)}{dz^n} \right|_{z=0} = \langle n \rangle^m$$

$$\Rightarrow \quad f_1 = \langle n \rangle \quad \text{and} \quad f_m \equiv 0, \text{ for } m > 1 \text{ - expected, since Poisson state}$$

 $f_1 = \langle n \rangle$  and  $f_m \equiv 0$ , for m > 1 - expected, since Poisson statistics implies production of uncorrelated particles and cumulants of order  $m \ge 2$  must vanish.

## Two-particle azimuthal correlations

- Energy-momentum conservation (e.g. resonances' decays,jets).
- Restricting the variables y<sub>1</sub> and y<sub>2</sub> to represent the azimuthal production angles φ<sub>1</sub> and φ<sub>2</sub>, we have:

$$C_2(\phi_1, \phi_2) = \rho_2(\phi_1, \phi_2) - \rho_1(\phi_1)\rho_1(\phi_2)$$
$$R_2(\phi_1, \phi_2) = \frac{\rho_2(\phi_1, \phi_2)}{\rho_1(\phi_1)\rho_1(\phi_2)} - 1$$

where the densities  $\rho_1(\phi_i)$  and  $\rho_2(\phi_1, \phi_2)$ are measured for specific ranges of  $p_T^{\min} \leq p_T \leq p_T^{\max}$  and  $\eta_{\min} \leq \eta \leq \eta_{\max}$ .

• In the absence of polarization or other discriminating direction, one expects that  $\rho_1(\phi_1) = \rho_1(\phi_2) \equiv \bar{\rho}_1$ , and  $C_2$  should depend on  $\Delta \phi = \phi_1 - \phi_2$ :

• 
$$\phi_1, \phi_2 \rightarrow \Delta \phi, \ \bar{\phi} = (\phi_1 + \phi_2)/2$$
  
 $C_2(\Delta \phi) = \rho_2(\Delta \phi) - \frac{1}{2\pi} \int_0^{2\pi} \bar{\rho}_1^2 d\bar{\phi} = \rho_2(\Delta \phi) - \rho_2(\Delta \phi)$ 





## Correlations from anisotropic flow

 Two-particle correlations may be very much influenced by collective effects as in collisions of heavy nuclei.

